INTRODUCTION TO ESTIMATION THEORY

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DEFINITION (CONVERGENCE IN DISTRIBUTION)

Let random variables X_n and X have distributions $F_n(\cdot)$ and $F(\cdot)$ respectively. Then X_n is said to converge in distribution to X, if

$$F_n \Rightarrow F$$

or equivalently,

$$\lim_{n \to \infty} \Pr[X_n \le x] = \Pr[X \le x]$$

for every x such that Pr[X = x] = 0.

Remark: We only require $F_n(\cdot)$ to converge at every continuous point!

CONVERGENCE IN DISTRIBUTION

EXAMPLE

Take $p_{n,k} = \frac{\lambda}{n}$

$$\mu\left\{k\right\} = \binom{n}{k} (\frac{\lambda}{n})^k (1-\frac{\lambda}{n})^(n-k), \text{ for } 0 \leq k \leq k$$

Then

 $\mu_n \Rightarrow \mathsf{Poisson}(\lambda).$

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 $X_n \xrightarrow{\mathcal{P}} X$ implies $X_n \Rightarrow X$

THEOREM

$$X_n \xrightarrow{\mathcal{P}} X \text{ implies } X_n \Rightarrow X$$

PROOF.

Obseve that
$$Pr[A \le a] - Pr[|A - B|] \le Pr[B \le a + b]$$
.
Then

$$Pr[X \le x - \epsilon] - Pr[|X_n - X| > \epsilon] \le Pr[X_n \le (x - \epsilon) + \epsilon] = Pr[X_n \le x],$$

and

$$Pr[X_n \le x] - Pr[|X_n - X| > \epsilon] \le Pr[X \le x + \epsilon].$$

Hence we get

$$Pr[X \le x - \epsilon] - Pr[|X_n - X| > \epsilon] \le Pr[X_n \le x]$$
$$\le Pr[X \le x + \epsilon] + Pr[|X_n - X| > \epsilon]$$

 $X_n \xrightarrow{\mathcal{P}} X$ IMPLIES $X_n \Rightarrow X$

CON'D.

which implies that

$$Pr[X \le x - \epsilon] \le \liminf Pr[X_n \le x]$$
$$\le \limsup Pr[X_n \le x]$$
$$\le Pr[X \le X + \epsilon]$$

Therefore, for each continuous point, we have

$$\lim_{n \to \infty} \Pr[X_n \le x] = \Pr[X \le x]$$

Remark: $X_n \Rightarrow X$ does not imply $X_n \xrightarrow{\mathcal{P}} X$

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Counterexample for $X_n \Rightarrow X$ implying $X_n \xrightarrow{\mathcal{P}} X$

EXAMPLE

 $X \perp Y$ and $X \sim Ber(p)$, $Y \sim Ber(p)$. Let $X_n = X$, for all n. Then we have

$$X_n \Rightarrow Y, \text{ but } X_n \not\to Y$$

Remark: Convergence in distribution gives NO information between the correlation of each random variables!

THEOREM (SKOROGOD'S THEOREM)

Suppose μ_n and μ are probability measures on $(\mathbb{R}, \mathcal{B})$, and $\mu_n \Rightarrow \mu$. Then there exist random variables Y_n and Y such that:

- 1. they are both defined on common probability space $(\Omega, \mathcal{F}, \mathcal{P})$
- 2. $Pr[Y_n \leq y] = \mu_n(-\infty, y]$, for all y
- 3. $Pr[Y \le y] = \mu(-\infty, y]$, for all y

4.
$$\lim_{n \to \infty} Y_n(\omega) = Y(\omega)$$
, for every ω in Ω

Remark: This implies that cdfs are sufficient; we do not need to rely on the inherited probability space.

Historical Aspects

- CLT concerns the situation that the limit distribution of the normalized sum is normal
- As an example, for i.i.d. zero-mean sequence $X_1, X_2, ...$

$$\frac{X_1 + \ldots + X_n}{n\mathbb{E}[X_i^2]} \Rightarrow N$$

where \boldsymbol{N} has standard normal distribution

• **Question:** What is the rare of convergence of normalized sum distribution to standard normal distrubution?

- The first convergence rate estimates in the CLT were obtained by A.M, Lyapounov in 1900-1901
- In the beginning of 1940s, the classic Berry-Esseen estimate came to the light

THEOREM (BERRY-ESSEEN THEOREM (I.I.D. CASE))

$$\sup_{x \in \mathbf{R}} |F_n(x) - \Phi(x)| \le C \frac{\beta_3}{\sigma^3 \sqrt{n}}$$

where

•
$$F_n$$
 is the cdf of $\frac{X_1 + \ldots + X_n}{n\mathbb{E}[X_i^2]}$

• Φ is the standard normal cdf

•
$$\beta_3 = \mathbb{E}[|X - \mathbb{E}X|^3]$$

•
$$\sigma^2 = \mathbb{E}[|X - \mathbb{E}X|^2]$$

• C is a universal constant, independent of n, F_n

THEOREM (BERRY-ESSEEN THEOREM (INDEPENDENT CASE))

$$\sup_{x \in \mathbf{R}} |Pr[\frac{X_1 + \dots + X_n}{s_n} \le a] - \Phi(x)| \le C \frac{r_n}{s_n^3}$$

where

• Φ is the standard normal cdf

•
$$r_n = \mathbb{E}[|(X_1 + ... + X_n) - \mathbb{E}(X_1 + ... + X_n)|^3]$$

- $s_n^2 = \mathbb{E}[|(X_1 + ... + X_n) \mathbb{E}(X_1 + ... + X_n)|^2]$
- C is a universal constant, independent of n, F_n

Remark:

- Lower bound of $C: C \ge \frac{\sqrt{10} + 3}{6\sqrt{2\pi}} \approx 0.40973$ (Esseen (1956))
- Upper bound of $C\text{:}~C \leq 0.4785$ (Tyurin (2010))

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